

The phase-space parameters of the brightest halo galaxies

Frank C. van den Bosch,^{1★} Simone M. Weinmann,¹ Xiaohu Yang,² H. J. Mo,²
Cheng Li³ and Y. P. Jing⁴

¹*Department of Physics, Swiss Federal Institute of Technology, ETH Hönggerberg, CH-8093, Zurich, Switzerland*

²*Department of Astronomy, University of Massachusetts, 710 North Pleasant Street, Amherst MA 01003-9305, USA*

³*Center for Astrophysics, University of Science and Technology of China, Hefei, 230026, China*

⁴*Shanghai Astronomical Observatory; the Partner Group of MPA, Nandan Road 80, Shanghai 200030, China*

Accepted 2005 May 27. Received 2005 May 11; in original form 2005 February 23

ABSTRACT

The brightest galaxy in a dark matter halo is expected to reside at rest at the centre of the halo. In this paper, we test this ‘Central Galaxy Paradigm’ (CGP) using group catalogues extracted from the Two-Degree Field Galaxy Redshift Survey (2dFGRS) and the Sloan Digital Sky Survey (SDSS). For each group, we compute a parameter \mathcal{R} , which is defined as the difference between the velocity of the brightest group galaxy and the average velocity of the other group members (hereafter satellites), normalized by the unbiased estimator of the velocity dispersion of the satellite galaxies. Because the redshift surveys suffer from incompleteness effects and the group selection criterion unavoidably selects interlopers, a proper comparison between data and model needs to take this into account. To this extent, we use detailed mock galaxy redshift surveys (MGRSs), which are analysed in exactly the same way as the data, thus allowing for a fair comparison. We show that the CGP is only consistent with the data in haloes with $M \lesssim 10^{13} h^{-1} M_{\odot}$, while in more massive haloes the data indicate a non-zero offset between the brightest galaxy and the satellites. This indicates that either central galaxies reside at the minimum of the dark matter potential, but that the halo itself is not yet fully relaxed, or, that the halo is relaxed, but that the central galaxy oscillates in its potential well. The former is consistent with the fact that the velocity bias of the brightest halo galaxies is larger in more massive haloes, while the latter may be indicative of cored, rather than cusped, dark matter haloes. We discuss several implications of these findings, including mass estimates based on satellite kinematics, strong gravitational lensing, halo occupation models, and the frequency and longevity of lopsidedness in disc galaxies.

Key words: methods: statistical – galaxies: haloes – galaxies: kinematics and dynamics – dark matter.

1 INTRODUCTION

In the standard picture of galaxy formation, hot gas in virialized dark matter haloes cools and accumulates at the centre of the potential well, where it forms a galaxy (White & Rees 1978). During the hierarchical build-up of larger and larger structures, haloes with their ‘central’ galaxies are accumulated by even larger haloes. At that point, the halo becomes a subhalo and the central galaxy becomes a satellite galaxy. In the standard picture, it is envisioned that a satellite galaxy no longer accretes hot gas, which instead is only accreted by the galaxy in the centre of the potential well (e.g. Kauffmann, White & Guiderdoni 1993; Somerville & Primack 1999; Cole et al. 2000). Because this central galaxy therefore con-

tinues to grow, it is expected to be the brightest, most massive galaxy in a halo. This is further assured by the fact that any other massive galaxy would quickly sink to the centre of the potential well by dynamical friction to merge with the central galaxy, thus producing an even more massive central galaxy. Therefore, according to the standard paradigm, the brightest galaxy in a halo will reside at rest at the centre of the potential well. Note that this is clearly a statistical statement, as it does not necessarily hold for each individual system (e.g. non-virialized, strongly interacting systems). Hereafter, we will refer to this paradigm as the ‘Central Galaxy Paradigm’ (CGP), and use the terms ‘central galaxy’ and ‘brightest halo galaxy’ without distinction.

The CGP plays an important role in various areas of astrophysics. For example, attempts to measure halo masses from the kinematics of satellite galaxies are always based on the general assumption that the ‘host’ galaxy is located at rest at the centre of a relaxed halo (e.g.

★E-mail: vdbosch@phys.ethz.ch

Zaritsky et al. 1993; McKay et al. 2002; Brainerd & Specian 2003; Prada et al. 2003; van den Bosch et al. 2004). This assumption is also used in virtually all mass models of strong gravitational lenses. On the other hand, the observed frequency and longevity of lopsidedness in disc galaxies (e.g. Richter & Sancisi 1994; Zaritsky & Rix 1997) is often interpreted as evidence for an actual offset between galaxy and halo (e.g. Levine & Sparke 1998). The CGP also plays a role in halo occupation modelling, where assumptions have to be made regarding the spatial distribution of galaxies in haloes in order to compute the galaxy–galaxy correlation function on small scales (e.g. Scoccimarro et al. 2001; Berlind & Weinberg 2002; van den Bosch, Yang & Mo 2003; Magliocchetti & Porciani 2003; Yang, Mo & van den Bosch 2003; Tinker et al. 2004; Zehavi et al. 2004; Zheng et al. 2004).

This special dynamical status of the brightest galaxy in a halo has been tested for the special class of central dominant (cD) galaxies. Jones et al. (1979) have shown that cDs are located at the peak of the cluster X-ray emission, while Quintana & Lawrie (1982) used the kinematics of cluster galaxies to argue that cDs are at rest with respect to the cluster. Although this is in agreement with the CGP, more recent studies have revealed various cases in which the cD galaxy has a significant peculiar velocity with respect to the mean velocity of the other cluster members (e.g. Hill et al. 1988; Sharples, Ellis & Gray 1988; Zabludoff, Huchra & Geller 1990; Oegerle & Hill 1994, 2001). Applying a similar study to a dozen poor groups, Mulchaey & Zabludoff (1998) and Zabludoff & Mulchaey (1998) found that the position of the brightest galaxy in each group is indistinguishable from that of the group centre or from the centre of the X-ray emission. To our knowledge, however, the CGP has never been tested for a statistically significant sample of dark matter haloes that span a wide range in masses. In this paper, we use data from the Two-Degree Field Galaxy Redshift Survey (2dFGRS; Colless et al. 2001) and the Sloan Digital Sky Survey (SDSS; York et al. 2000) to directly test whether the brightest galaxies in dark matter haloes are located at rest at the centre of their potential well. We show that, although the brightest halo galaxies are clearly segregated with respect to the other galaxies in the same halo, they have a non-zero specific kinetic energy, at least in haloes with $M \gtrsim 10^{13} h^{-1} M_{\odot}$.

This paper is organized as follows. In Section 2, we present a statistic that can be used to test the CGP, which we apply to the 2dFGRS and SDSS in Section 3. In Section 4, we describe a simple model for the velocity and spatial bias of the brightest halo galaxies, which we use in Section 5 to construct detailed mock galaxy redshift surveys (hereafter MGRSs) of the 2dFGRS. In Section 6, we compare these mock-ups with the data in order to constrain the phase-space parameters of the brightest halo galaxies. Section 7 discusses various implications of our results and we summarize our conclusions in Section 8.

2 DYNAMICAL SIGNATURE OF CENTRAL GALAXIES

Observationally, the only kinematic information that is available to test the CGP is the line-of-sight velocities obtained from redshifts. In what follows, we use v_c to refer to the line-of-sight velocity of the brightest halo galaxy and v_i is the line-of-sight velocity of the i th satellite galaxy. In addition, we define the difference $\Delta V = \bar{v}_s - v_c$ between the *mean* velocity of the satellite galaxies

$$\left(\bar{v}_s = \frac{1}{N_s} \sum_{i=1}^{N_s} v_i \right)$$

and that of the central galaxy. If the CGP is correct and v_i follows a Gaussian distribution with velocity dispersion σ_s , the probability that a halo with N_s satellite galaxies has a value of ΔV is given by

$$P(\Delta V)d\Delta V = \frac{1}{\sqrt{2\pi}\sigma} \exp \left[-\frac{(\Delta V)^2}{2\sigma^2} \right] d\Delta V, \quad (1)$$

with $\sigma = \sigma_s/\sqrt{N_s}$. Therefore, in principle, one could define the parameter

$$R = \frac{\sqrt{N_s}(\bar{v}_s - v_c)}{\sigma_s}, \quad (2)$$

and test the CGP by checking whether R follows a normal distribution with zero mean and unit variance. However, the velocity dispersion σ_s is generally unknown and we have to use its unbiased estimator

$$\hat{\sigma}_s = \sqrt{\frac{1}{N_s - 1} \sum_{i=1}^{N_s} (v_i - \bar{v}_s)^2} \quad (3)$$

instead. This allows us to define the modified parameter

$$\mathcal{R} = \frac{\sqrt{N_s}(\bar{v}_s - v_c)}{\hat{\sigma}_s}. \quad (4)$$

If the null hypothesis of the CGP is correct, \mathcal{R} should follow a Student t distribution with $\nu = N_s - 1$ degrees of freedom. Note that $P_\nu(\mathcal{R})$ approaches a normal distribution with zero mean and unit variance in the limit $N_s \rightarrow \infty$.

The applicability of this ‘ \mathcal{R} test’ is strongly related to the ability to find those galaxies that belong to the same dark matter halo. To this extent, we use the halo-based galaxy group finder developed by Yang et al. (2005a, hereafter YMBJ), which has been optimized for this task. Although this group finder is well tested and calibrated, it is not perfect. In particular, because of redshift errors and redshift space distortions, it is unavoidable that one selects interlopers (galaxies that are not associated with the same halo). The expectation value of $|v_s - v_c|$ will be larger for an interloper than for a true satellite. As long as the interloper is fainter than the brightest galaxy in the group (halo) to which it is assigned, its impact on \mathcal{R} may be small, as it affects both the numerator and the denominator. However, if the interloper is brighter than all true group members, $|\mathcal{R}|$ will typically be severely overestimated. Another problem is related to the fact that the 2dFGRS and SDSS suffer from various incompleteness effects. If the actual brightest halo galaxy is missed (i.e. is not present in the survey), \mathcal{R} will be measured with respect to a satellite galaxy, which again will bias $|\mathcal{R}|$ high. The presence of interlopers and incompleteness effects, therefore, tends to create excessive wings in the \mathcal{R} distribution. A comparison with the Student t distribution might then give the wrong impression that the null hypothesis is rejected. Because the typical occupation numbers of haloes are small, this effect can be very strong, as we demonstrate in Section 6. To circumvent these problems, we compare the \mathcal{R} distributions obtained from groups in the 2dFGRS and SDSS against those obtained from groups extracted from detailed MGRSs, which suffer from interlopers and incompleteness effects to the same extent as the real data.

3 APPLICATION TO THE 2DFGRS AND SDSS

3.1 Group selection

The \mathcal{R} test described above requires a selection of galaxies that belong to the same dark matter halo. In YMBJ, we developed a halo-based galaxy group finder, that is optimized for this task. Here, we

give a brief description of this group finder and refer the interested reader to YMBJ for details.

The basic idea behind our group finder is similar to that of the matched filter algorithm developed by Postman et al. (1996), although it also makes use of the galaxy kinematics. The group finder starts with an assumed mass-to-light ratio to assign a tentative mass to each potential group, identified using the friends-of-friends (FOF) method. This mass is used to estimate the size and velocity dispersion of the underlying halo that hosts the group, which in turn is used to determine group membership (in redshift space). This procedure is iterated until no further changes occur in group memberships. Using detailed MGRSs, the performance of our group finder has been tested in terms of completeness of true members and contamination by interlopers. The average completeness of individual groups is ~ 90 per cent and with only ~ 20 per cent interlopers. Furthermore, the resulting group catalogue is insensitive to the initial assumption regarding the mass-to-light ratios and is more successful than the conventional FOF method in associating galaxies according to their common dark matter haloes.

3.2 The 2dFGRS

We use the final, public data release from the 2dFGRS, restricting ourselves only to galaxies with redshifts $0.01 \leq z \leq 0.20$ in the North Galactic Pole (NGP) and South Galactic Pole (SGP) subsamples with a redshift quality parameter $q \geq 3$ and a redshift completeness $c > 0.8$. This leaves a grand total of 151 280 galaxies with a sky coverage of $\sim 1125 \text{ deg}^2$. The typical rms redshift and magnitude errors are 85 km s^{-1} and 0.15 mag , respectively (Colless et al. 2001). Absolute magnitudes for galaxies in the 2dFGRS are computed using the K -corrections of Madgwick et al. (2002).

Application of the halo-based group finder to this galaxy sample yields a group catalogue consisting of 77 708 systems. Detailed information regarding the clustering properties and galaxy occupation statistics of these groups can be found in YMBJ and Yang et al. (2005b,c). In what follows, we restrict our analyses to the 2502 groups in this catalogue with four members or more.

3.3 The SDSS

In addition to the 2dFGRS, we also use data from the SDSS. In particular, we use the New York University Value-Added Galaxy Catalogue (NYU-VAGC),¹ described in detail in Blanton et al. (2005). The NYU-VAGC is based on the SDSS Data Release 2 (Abazajian et al. 2004), but with an independent set of significantly improved reductions. From this catalogue, we select all galaxies in the Main Galaxy Sample, which has an extinction corrected Petrosian magnitude limit of $r = 18$. We prune this sample to those galaxies in the redshift range $0.01 \leq z \leq 0.20$ and with a redshift completeness $c > 0.7$. This leaves a grand total of 184 425 galaxies with a sky coverage of $\sim 1950 \text{ deg}^2$. From this SDSS sample, we construct a group catalogue that contains 102 935 systems. A more detailed description of this catalogue will be presented in Weinmann et al. (in preparation). As for the 2dFGRS, we restrict our analysis to the groups with four members or more, of which there are 3473 in our catalogue.

3.4 Comparison of 2dFGRS with SDSS

For each group in both the 2dFGRS and SDSS catalogues described above, we compute \mathcal{R} . Fig. 1 plots the cumulative distributions of $|\mathcal{R}|$ for both surveys. In the upper-left panel, we plot the distributions using all groups in the range $50 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 1000 \text{ km s}^{-1}$, with $\hat{\sigma}_s$ the unbiased estimator of σ_s (equation 3). In the other three panels, we plot $P(<|\mathcal{R}|)$ for three subsamples (the values in square brackets indicate the range in $\hat{\sigma}_s$ used, in km s^{-1}). Overall, the agreement between SDSS and 2dFGRS is extremely good. To make the comparison more quantitative, we use the Kolmogorov–Smirnov (hereafter KS) test to compute the probability P_{KS} that both $P(|\mathcal{R}|)$ are drawn from the same distribution. The resulting probabilities are indicated in each panel. These confirm what can already be inferred by eye, namely that both \mathcal{R} distributions are consistent with each other. Given this good agreement between both data sets, we only concentrate on the 2dFGRS in what follows. The main reason for choosing this survey over the SDSS is that we have accurate mock-ups for the 2dFGRS that have been well tested. Given the good agreement between 2dFGRS and SDSS, we argue that any result based on the former will also hold for the latter.

4 MODELLING VELOCITY BIAS OF CENTRAL GALAXIES

The main goal of this paper is to use the \mathcal{R} distributions presented above in order to constrain the phase-space parameters of the brightest halo galaxies. We will express these in terms of their spatial and velocity bias with respect to the satellites. If the null hypothesis of the CGP is correct, both the spatial and the velocity bias should equal zero. In order to model these biases and to incorporate them in the mock redshift surveys that we will use for comparison with the data, we proceed as follows.

We assume that each dark matter halo has an NFW (Navarro, Frenk & White 1997) density distribution, $\rho_{\text{dm}}(r)$, with virial radius r_{vir} , characteristic scale radius r_s , and concentration parameter $c = r_{\text{vir}}/r_s$. Assuming haloes to be spherical and isotropic, the local, one-dimensional velocity dispersion follows from solving the Jeans equation

$$\sigma_{\text{dm}}^2(r) = \frac{1}{\rho_{\text{dm}}(r)} \int_r^\infty \rho_{\text{dm}}(r') \frac{\partial \Psi}{\partial r}(r') dr', \quad (5)$$

with $\Psi(r)$ the gravitational potential (Binney & Tremaine 1987). Using that $\Psi/r = GM(r)/r^2$ and defining the virial velocity $V_{\text{vir}} = \sqrt{GM/r_{\text{vir}}}$, we obtain

$$\sigma_{\text{dm}}^2(r) = V_{\text{vir}}^2 \frac{c}{f(c)} \left(\frac{r}{r_s} \right) \left(1 + \frac{r}{r_s} \right)^2 \mathcal{I}(r/r_s), \quad (6)$$

with $f(x) = \ln(1+x) - x/(1+x)$ and

$$\mathcal{I}(y) = \int_y^\infty \frac{f(\tau) d\tau}{\tau^3(1+\tau)^2}. \quad (7)$$

The halo-averaged velocity dispersion is given by

$$\begin{aligned} \langle \sigma_{\text{dm}} \rangle_{\text{M}} &\equiv \frac{4\pi}{M} \int_0^{r_{\text{vir}}} \rho_{\text{dm}}(r) \sigma_{\text{dm}}(r) r^2 dr \\ &= V_{\text{vir}} \sqrt{\frac{c}{f^3(c)}} \int_0^c \frac{y^{3/2} \mathcal{I}^{1/2}(y)}{(1+y)} dy \end{aligned} \quad (8)$$

(cf. van den Bosch et al. 2004).

Throughout this paper, we assume that the N_{sat} satellite galaxies in a halo of mass M follow a number density distribution

¹ <http://wassup.physics.nyu.edu/vagc/#download>

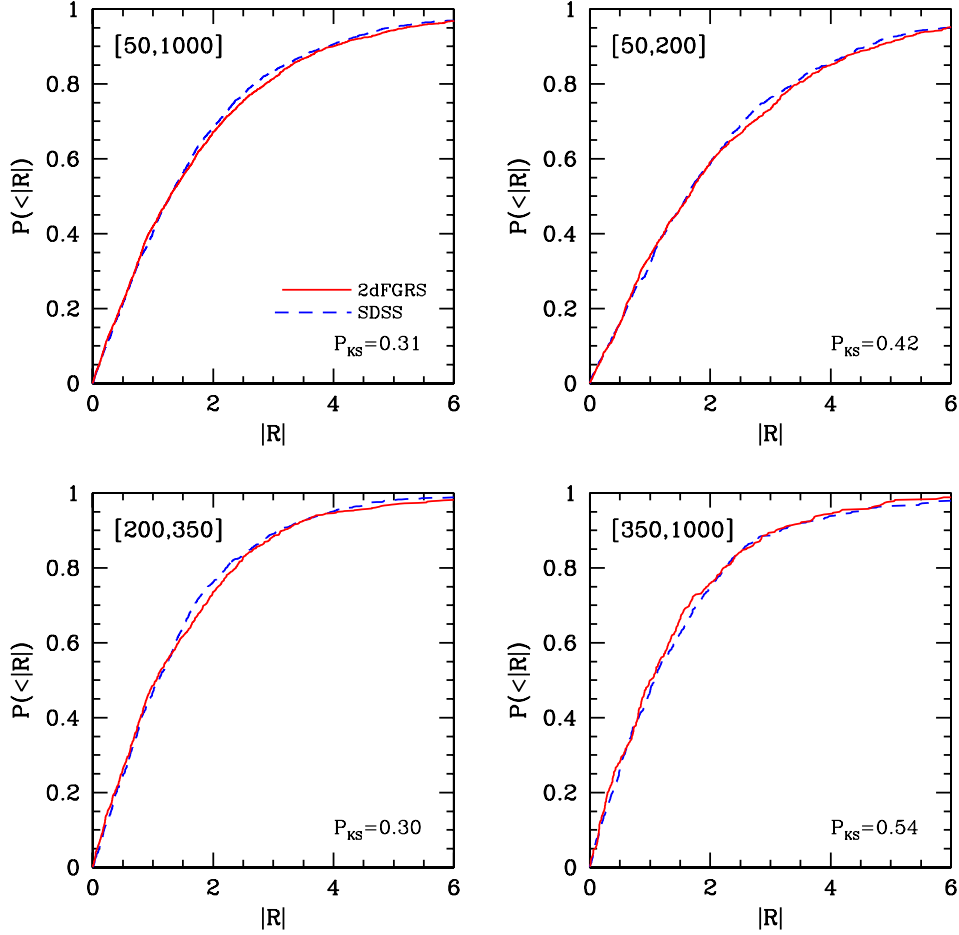


Figure 1. A comparison of the cumulative distributions of $|\mathcal{R}|$ of 2dFGRS [solid lines (red online)] and SDSS [dashed lines (blue)] groups. Results are shown for four intervals in $\hat{\sigma}_s$, indicated in square brackets in each panel. The KS probability, P_{KS} , that both distributions are drawn from the same distribution is also indicated. Note that the $P(|\mathcal{R}|)$ from 2dFGRS and SDSS are in excellent agreement with each other.

$n_{\text{sat}}(r) = (N_{\text{sat}}/M)\rho_{\text{dm}}(r)$, i.e. there is no spatial bias between satellite galaxies and dark matter particles. As shown in van den Bosch et al. (2005), this is consistent with the observed radial distribution of satellite galaxies in the 2dFGRS. If we further assume that the satellites are in isotropic equilibrium, it also follows that there is no velocity bias between the satellites and the dark matter, neither globally (i.e. $\langle\sigma_{\text{sat}}\rangle_{\text{M}} = \langle\sigma_{\text{dm}}\rangle_{\text{M}}$) nor locally [i.e. $\sigma_{\text{sat}}(r) = \sigma_{\text{dm}}(r)$].

When stacking all haloes of a given mass, we assume that their brightest halo galaxies follow a number density distribution given by a Hernquist (1990) profile:²

$$\rho_{\text{cen}}(r) \propto \frac{1}{2\pi} \frac{a}{r} \frac{1}{(r+a)^3}. \quad (9)$$

This implies a probability distribution for r of the central galaxies of

$$P_{\text{cen}}(r)dr = 2 \left(\frac{r_{\text{vir}} + a}{r_{\text{vir}}} \right)^2 \frac{ar}{(r+a)^3} dr. \quad (10)$$

In order to parametrize the characteristic radius a in terms of that of the dark matter halo, we define the parameter $f_{\text{cen}} \equiv a/r_s$. A brightest halo galaxy at a halo-centric radius r has an isotropic velocity

dispersion

$$\begin{aligned} \sigma_{\text{cen}}^2(r) &= \frac{1}{\rho_{\text{cen}}(r)} \int_r^\infty \rho_{\text{cen}}(r') \frac{\partial \Psi}{\partial r}(r') dr' \\ &= V_{\text{vir}}^2 \frac{c}{f(c)} \left(\frac{r}{r_s} \right) \left(f_{\text{cen}} + \frac{r}{r_s} \right)^3 \mathcal{J}(r/r_s), \end{aligned} \quad (11)$$

with

$$\mathcal{J}(y) = \int_y^\infty \frac{f(\tau) d\tau}{\tau^3 (f_{\text{cen}} + \tau)^3}. \quad (12)$$

This implies a halo-averaged velocity dispersion of

$$\begin{aligned} \langle\sigma_{\text{cen}}\rangle_{\text{M}} &\equiv \frac{\int_0^{r_{\text{vir}}} \rho_{\text{cen}}(r) \sigma_{\text{cen}}(r) r^2 dr}{\int_0^{r_{\text{vir}}} \rho_{\text{cen}}(r) r^2 dr} \\ &= V_{\text{vir}} \sqrt{\frac{4c}{f(c)}} f_{\text{cen}} \int_0^c \frac{y^{3/2} \mathcal{J}^{1/2}(y)}{(f_{\text{cen}} + y)^{3/2}} dy, \end{aligned} \quad (13)$$

which allows us to define the velocity bias of the brightest halo galaxies as $b_{\text{vel}} \equiv \langle\sigma_{\text{cen}}\rangle/\langle\sigma_{\text{dm}}\rangle = \langle\sigma_{\text{cen}}\rangle/\langle\sigma_{\text{sat}}\rangle$. In addition to the velocity bias, we define the spatial bias as $b_{\text{rad}} \equiv \langle r_{\text{cen}} \rangle / \langle r_{\text{dm}} \rangle = \langle r_{\text{cen}} \rangle / \langle r_{\text{sat}} \rangle$, where the expectation value for the radius follows from

$$\langle r \rangle = \frac{\int_0^{r_{\text{vir}}} \rho(r) r^3 dr}{\int_0^{r_{\text{vir}}} \rho(r) r^2 dr}. \quad (14)$$

² The choice for this particular distribution is not motivated by any physical considerations, other than the fact that it is well behaved, both at $r = 0$ and at $r \rightarrow \infty$.

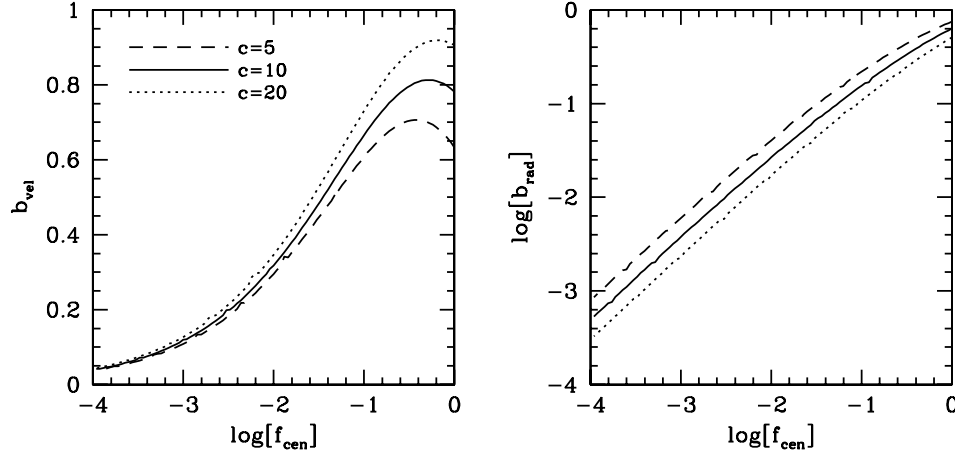


Figure 2. The velocity bias (left-hand panel) and spatial bias (right-hand panel) of central galaxies as function of the parameter f_{cen} , which expresses the characteristic scale of the radial distribution of central galaxies in terms of the characteristic scale of the NFW density distribution (see Section 4). Results are shown for three values of the halo concentration parameter c , as indicated.

For an NFW density distribution with concentration c , this reduces to

$$\langle r_{\text{dm}} \rangle = \left[\frac{(2+c)/(1+c) - (2/c)\ln(1+c)}{\ln(1+c) - c/(1+c)} \right] r_{\text{vir}} \quad (15)$$

or $\langle r_{\text{dm}} \rangle = 0.41 r_{\text{vir}}$ for $c = 10$.

Fig. 2 plots b_{vel} (left-hand panel) and b_{rad} (right-hand panel) as function of f_{cen} for three values of the halo concentration parameter c . In the limit $f_{\text{cen}} \rightarrow 0$, the probability distribution $P_{\text{cen}}(r)$ becomes a Dirac delta function. This implies that the central galaxy is sitting still at the centre of the dark matter halo (i.e. the null hypothesis of the CGP), so that $b_{\text{vel}} = b_{\text{rad}} = 0$. Increasing f_{cen} increases the probability to find the brightest halo galaxy at larger halo-centric radii, which corresponds to a larger velocity bias. Note, however, that b_{vel} never approaches unity, which is due to the fact that $\rho_{\text{cen}}(r)$ can not be made to match $\rho_{\text{dm}}(r)$ for any value of f_{cen} . Typically $b_{\text{vel}} \gg b_{\text{rad}}$, which is a reflection of the ‘depth’ of the NFW potential. For example, for a velocity bias of $b_{\text{vel}} = 0.5$ (i.e. corresponding to a specific kinetic energy that is one quarter of that of the satellites), the radial bias is $b_{\text{rad}} \simeq 0.07$ (assuming $c = 10$). Combining this with equation (15) implies an expectation value for the offset of the central galaxy from the dark matter distribution of $\langle r_{\text{cen}} \rangle \simeq 0.03 r_{\text{vir}}$. For a Milky Way sized system, this corresponds to ~ 5 kpc, comparable to the characteristic radius (scalelength) of the galaxy itself. This clearly demonstrates that the signal of an off-centred brightest halo galaxy is far more pronounced and thus easier to detect in velocity space than in real space (i.e. by measuring the projected distance between the brightest galaxy and the centroid of the satellites).

In what follows, we construct a set of MGRSs for different values of b_{vel} and compare the \mathcal{R} distributions of their groups against those of the 2dFGRS, which are statistically identical to those of the SDSS, in an attempt to constrain b_{vel} . Although the construction of these MGRSs is based on the model described above and thus depends on the assumed functional form for $P_{\text{cen}}(r)$, this does not impact on our results. The statistic that we use to constrain b_{vel} depends only on the velocities of the central galaxies; were we to adopt another $P_{\text{cen}}(r)$ but with the same value of b_{vel} , the resulting \mathcal{R} distributions would be virtually indistinguishable. The assumed form of $P_{\text{cen}}(r)$ only impacts on the relation between b_{vel} and b_{rad} .

5 MOCK GALAXY REDSHIFT SURVEYS

We construct MGRSs by populating dark matter haloes with galaxies of different luminosities. The distribution of dark matter haloes is obtained from a set of large N -body simulations (dark matter only) for a Λ cold dark matter (CDM) ‘concordance’ cosmology with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $h = 0.7$ and $\sigma_8 = 0.9$. In this paper, we use two simulations with $N = 512^3$ particles each, which are described in more detail in Jing & Suto (2002). The simulations have periodic boundary conditions and box sizes of $L_{\text{box}} = 100 h^{-1}$ Mpc (hereafter L_{100}) and $L_{\text{box}} = 300 h^{-1}$ Mpc (hereafter L_{300}). We follow Yang et al. (2004) and replicate the L_{300} box on a $4 \times 4 \times 4$ grid. The central $2 \times 2 \times 2$ boxes are replaced by a stack of $6 \times 6 \times 6$ L_{100} boxes and the virtual observer is placed at the centre (see fig. 11 in Yang et al. 2004). This stacking geometry circumvents incompleteness problems in the mock survey due to insufficient mass resolution of the L_{300} simulations and allows us to reach the desired depth of $z_{\text{max}} = 0.20$ in all directions.

Dark matter haloes are identified using the standard FOF algorithm with a linking length of 0.2 times the mean interparticle separation. Unbound haloes and haloes with less than 10 particles are removed from the sample. In Yang et al. (2004), we have shown that the resulting halo mass functions are in excellent agreement with the analytical halo mass function of Sheth, Mo & Tormen (2001).

5.1 Populating haloes with galaxies

In order to populate the dark matter haloes with galaxies of different luminosities, we use the conditional luminosity function (hereafter CLF), $\Phi(L|M)$, which gives the average number of galaxies of luminosity L that resides in a halo of mass M . As demonstrated in van den Bosch et al. (2003) and Yang et al. (2003), the CLF is well constrained by the galaxy luminosity function and by the galaxy–galaxy correlation lengths as functions of luminosity. In the MGRSs used here, we use the CLF with ID no. 6 given in table 1 of van den Bosch et al. (2005). We have tested that none of our results depends significantly on this particular choice for the CLF.

Because of the mass resolution of the simulations and because of the completeness limit of the 2dFGRS, we adopt a minimum galaxy luminosity of $L_{\text{min}} = 10^7 h^{-2} L_\odot$. The *mean* number of galaxies

with $L \geq L_{\min}$ that resides in a halo of mass M is given by

$$\langle N \rangle_M = \int_{L_{\min}}^{\infty} \Phi(L|M) dL. \quad (16)$$

In order to Monte Carlo sample occupation numbers for individual haloes, one requires the full probability distribution $P(N|M)$ (with N an integer) of which $\langle N \rangle_M$ gives the mean. We differentiate between satellite galaxies and central galaxies. The total number of galaxies per halo is the sum of N_{cen} , the number of central galaxies, which is either one or zero, and N_{sat} , the (unlimited) number of satellite galaxies. We assume that N_{sat} follows a Poisson distribution and require that $N_{\text{sat}} = 0$ whenever $N_{\text{cen}} = 0$. The halo occupation distribution is thus specified as follows: if $\langle N \rangle_M \leq 1$ then $N_{\text{sat}} = 0$ and N_{cen} is either zero (with probability $P = 1 - \langle N \rangle_M$) or one (with probability $P = \langle N \rangle_M$). If $\langle N \rangle_M > 1$ then $N_{\text{cen}} = 1$ and N_{sat} is drawn from a Poisson distribution with a mean of $\langle N \rangle_M - 1$.

We follow Yang et al. (2004) and draw the luminosity of the brightest galaxy in each halo from $\Phi(L|M)$ using the restriction that $L > L_1$ with L_1 defined by

$$\int_{L_1}^{\infty} \Phi(L|M) dL = 1. \quad (17)$$

The luminosities of the satellite galaxies are also drawn from $\Phi(L|M)$, but with the restriction $L_{\min} < L < L_1$.

Next, we assign all galaxies a position and velocity within their halo, using the number density distributions and (isotropic) velocity dispersion profiles given in Section 4. Note that this implicitly assumes that all haloes, as well as their galaxy populations, are relaxed. Halo concentrations as function of halo mass are computed using the relation given by Eke, Navarro & Steinmetz (2001).

5.2 Creating mock surveys

The 2dFGRS uses a multifibre spectrograph to obtain redshifts. However, because of the physical size of the fibers, when two galaxies are closer than ~ 30 arcsec in projection only one of them can be targeted. Furthermore, due to clustering, some areas on the sky contain more galaxies within a single two-degree field than the available number of fibers. By using a sophisticated tiling strategy, these problems are largely overcome, yielding a fairly uniform sampling rate. Nevertheless, some spatial non-uniformities remain. In addition, fainter galaxies yield noisier spectra and therefore less accurate redshifts. All these effects combined result in a redshift completeness that depends on both position on the sky and on apparent magnitude. The 2dFGRS team has constructed maps that parametrize this position and magnitude dependent completeness (Colless et al. 2001; Norberg et al. 2002), and that facilitate a simulation of these effects in our MGRSs. However, as it turns out, the completeness depends also on the angular separation, θ , between galaxy *pairs* (see Hawkins et al. 2003). This is largely due to the problem of fiber collisions, which has not been completely corrected for by the tiling strategy. Finally, Norberg et al. (2002) have shown that the *parent* catalogue of the 2dFGRS, the APM catalogue, is only 91 per cent complete. As shown in van den Bosch et al. (2005), this incompleteness is, at least partially, due to image blending in the APM catalogue (see also Cole et al. 2001). Based on this information, we mimic the various observational selection and completeness effects in the 2dFGRS using the following steps.

(i) We define an (α, δ) -coordinate frame with respect to the virtual observer at the centre of the stack of simulation boxes, and remove all galaxies that are not located in the areas equivalent to the NGP and SGP regions of the 2dFGRS.

(ii) For each galaxy, we compute the apparent magnitude according to its luminosity and distance, to which we add an rms error of 0.15 mag. Because galaxies in the 2dFGRS were pruned by apparent magnitude *before* a K -correction was applied, we proceed as follows: we first apply a negative K -correction, then select galaxies according to the position-dependent magnitude limit (obtained using the apparent magnitude limit masks provided by the 2dFGRS team) and finally K -correct the magnitudes back to their rest-frame b_J band. Throughout, we use the type-dependent K -corrections given in Madgwick et al. (2002).

(iii) For each galaxy, we compute the redshift as ‘seen’ by the virtual observer. We take the observational velocity uncertainties into account by adding a random velocity drawn from a Gaussian distribution with dispersion 85 km s^{-1} .

(iv) To take account of the position- and magnitude-dependent completeness of the 2dFGRS, we randomly sample each galaxy using the completeness masks provided by the 2dFGRS team.

(v) To take account of the fiber-collision induced incompleteness, we compute the angular separations θ between all galaxy pairs and remove galaxies based on a probability $p(\theta)$, which we tune (by trial and error) so that we reproduce the pair-separation incompleteness quantified by Hawkins et al. (2003).

(vi) To take account of the incompleteness in the APM catalogue due to image blending, we model the characteristic size of a galaxy as

$$R_{\text{gal}} = 15 h^{-1} \text{ kpc} \left(\frac{L}{10^{10} h^{-2} L_{\odot}} \right)^{1/3} \quad (18)$$

and define the critical projection angle $\theta_{\text{max}} = R_{\text{gal}}/D_A$, with D_A the angular diameter distance of the galaxy. We then remove the faintest galaxy from all pairs for which $\theta < \theta_{\text{max}}$.

(vii) Finally, we remove a number of galaxies completely at random to bring the total fraction of removed galaxies, including those removed under (v) and (vi), to 9 per cent.

As shown in van den Bosch et al. (2005), this procedure results in mock 2dFGRS catalogues that accurately mimic all the various incompleteness effects, allowing for a direct, one-to-one comparison with the true 2dFGRS.

6 RESULTS

Using the method outlined above, we construct a set of ten MGRSs that only differ in the value of b_{vel} . Table 1 lists these mock-ups together with their corresponding values of b_{rad} and f_{cen} , computed for a halo with a concentration parameter $c = 10$. Note that these values of b_{rad} and f_{cen} are only to give the reader an order of magnitude estimate of the corresponding radial bias. The \mathcal{R} test used here only constrains the value of b_{vel} ; the corresponding values of b_{rad} and f_{cen} can only be computed indirectly and are model dependent (see Section 4).

In the case of $M_{1.0}$, we deviated somewhat from the procedure described in Section 4. Rather than giving the central galaxies a probability distribution (10), we simply treat the brightest galaxy as a satellite galaxy so that $b_{\text{vel}} = 1.0$. Note that, in this case, f_{cen} is not defined. For each of our ten MGRSs, we construct a group catalogue as described in Section 3.1, using those mock galaxies that are in the redshift range $0.01 \leq z \leq 0.20$ and with a completeness $c > 0.8$ (this mimics our selection from the 2dFGRS). In what follows, we restrict our analysis to groups with four or more members and with $50 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 1000 \text{ km s}^{-1}$.

Table 1. Comparison between MGRSs and 2dFGRS.

MGRS (1)	b_{vel} (2)	b_{rad} (3)	f_{cen} (4)	$P_{\text{KS}}[50, 1000]$ (5)	$P_{\text{KS}}[50, 200]$ (6)	$P_{\text{KS}}[200, 350]$ (7)	$P_{\text{KS}}[350, 1000]$ (8)
$M_{0.0}$	0.0	0.0	0.0	1.7×10^{-1}	2.6×10^{-1}	5.1×10^{-4}	2.9×10^{-2}
$M_{0.1}$	0.1	2.8×10^{-3}	7.3×10^{-4}	1.4×10^{-1}	8.7×10^{-1}	2.3×10^{-1}	7.4×10^{-2}
$M_{0.2}$	0.2	1.0×10^{-2}	3.3×10^{-3}	1.3×10^{-1}	4.9×10^{-1}	6.2×10^{-1}	8.7×10^{-2}
$M_{0.3}$	0.3	2.4×10^{-2}	8.5×10^{-3}	2.0×10^{-1}	4.3×10^{-1}	2.4×10^{-1}	1.6×10^{-1}
$M_{0.4}$	0.4	4.2×10^{-2}	1.8×10^{-2}	5.2×10^{-6}	8.2×10^{-2}	1.3×10^{-1}	3.5×10^{-1}
$M_{0.5}$	0.5	7.2×10^{-2}	3.5×10^{-2}	1.8×10^{-6}	7.5×10^{-2}	1.0×10^{-1}	3.8×10^{-1}
$M_{0.6}$	0.6	1.1×10^{-1}	6.6×10^{-2}	2.5×10^{-9}	4.0×10^{-2}	1.1×10^{-3}	1.2×10^{-1}
$M_{0.7}$	0.7	1.8×10^{-1}	1.3×10^{-1}	1.7×10^{-10}	3.3×10^{-3}	1.2×10^{-2}	7.0×10^{-2}
$M_{0.8}$	0.8	3.4×10^{-1}	3.3×10^{-1}	2.5×10^{-14}	2.4×10^{-5}	1.1×10^{-3}	7.4×10^{-3}
$M_{1.0}$	1.0	1.0	—	2.8×10^{-18}	1.9×10^{-5}	1.3×10^{-6}	5.3×10^{-6}

The MGRSs used for comparison with the 2dFGRS. Column (1) lists the ID of the MGRS. Columns (2), (3) and (4) list the velocity bias, spatial bias and value of f_{cen} , respectively (see Section 4 for definitions). Finally, columns (5)–(8) list the KS probabilities P_{KS} that the distributions of \mathcal{R} extracted from these MGRSs are consistent with those of the 2dFGRS for four different intervals in $\hat{\sigma}_s$, indicated by the values in square brackets (in km s^{-1}).

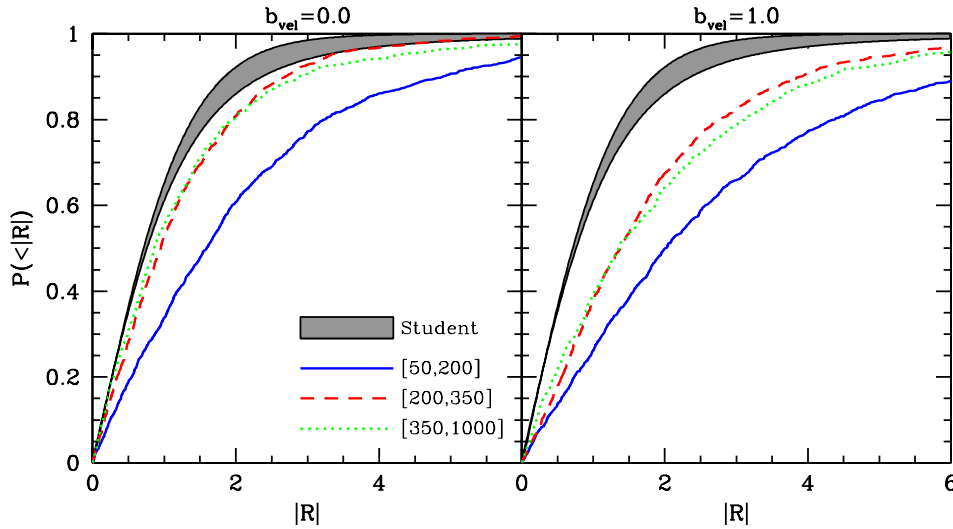


Figure 3. The cumulative distributions of $|\mathcal{R}|$ obtained from MGRSs $M_{0.0}$ (left-hand panel) and $M_{1.0}$ (right-hand panel), in which the brightest halo galaxies have a velocity bias of $b_{\text{vel}} = 0$ and $b_{\text{vel}} = 1$, respectively. Solid, dashed and dotted curves correspond to group samples with $50 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 200 \text{ km s}^{-1}$, $200 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 350 \text{ km s}^{-1}$, and $350 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 1000 \text{ km s}^{-1}$, respectively. The grey area indicates the area bounded by Student distributions with 3 and 9 degrees of freedom. In the ideal case without interlopers, the $P(< |\mathcal{R}|)$ of $M_{0.0}$ should fall in this range. The fact that they do not illustrates the impact of interlopers and completeness effects, and emphasizes the importance of using MGRSs for a fair comparison with the data. Finally, the fact that the $P(< |\mathcal{R}|)$ of $M_{0.0}$ and $M_{1.0}$ are significantly different illustrates that the \mathcal{R} test does have the ability to constrain the phase-space parameters of the brightest halo galaxies.

To illustrate the importance of using MGRSs, the left-hand panel of Fig. 3 shows the cumulative distributions of $|\mathcal{R}|$ obtained from the groups in $M_{0.0}$ for which $b_{\text{vel}} = 0.0$. In this mock-up, all the brightest halo galaxies have been located at rest at the centre of the halo. The grey area indicates the area bounded by Student t distributions with 3 and 9 degrees of freedom (corresponding to systems with four and 10 satellites, respectively), which spans the range covered by the vast majority of our groups). In principle, because this MGRS obeys the null hypothesis of the CGP, the resulting $P(< |\mathcal{R}|)$ should fall in this range. Clearly it does not, especially not for groups with $50 \text{ km s}^{-1} \leq \hat{\sigma}_s \leq 200 \text{ km s}^{-1}$ (solid line). This owes to the completeness effects in the survey, and to the fact that our group finder is not perfect and (unavoidably) selects interlopers. As we discussed in Section 2, these effects systematically broaden $P(\mathcal{R})$. Because the impact of one or two interlopers is much stronger in low-mass groups, which have fewer members, the $P(\mathcal{R})$ of groups with low $\hat{\sigma}_s$ deviates more from the predicted Student t distribution than that of more massive

groups. This clearly demonstrates that one needs to take interlopers and completeness effects into account, in a statistical sense, when interpreting the distribution of \mathcal{R} obtained from the 2dFGRS. The MGRSs used here are ideally suited for this task.

The right-hand panel of Fig. 3 shows the same results as in the left-hand panel, but now based on $M_{1.0}$ for which $b_{\text{vel}} = 1.0$. Clearly, for this MGRS the $P(\mathcal{R})$ are significantly broader than for $M_{0.0}$. This demonstrates that, despite the interloper/completeness problem, the detailed distributions of \mathcal{R} obtained from group catalogues do contain useful information that we can use to constrain b_{vel} .

Before we compare the \mathcal{R} distributions obtained from our MGRSs with those obtained from the 2dFGRS, we need to address two specific issues. The first is cosmic variance. Although the 2dFGRS is a large redshift survey, cosmic variance may still play a role. Because the \mathcal{R} distributions depend only mildly on b_{vel} , small differences in $P(\mathcal{R})$ due to cosmic variance may have a non-negligible impact on the inferred value of b_{vel} . In order to address this, we have

Table 2. The KS probabilities that express how similar the $P(\mathcal{R})$ distributions are for four independent mock-ups (all with $b_{\text{vel}} = 0.5$). Column (1) lists the numbers (arbitrary) of the mocks that are compared. Columns (2)–(5) list the KS probabilities that the \mathcal{R} distributions for groups in these mock-ups are consistent with being drawn from the same distribution. The different columns correspond to different σ_8 intervals as indicated at the top of each column. Note that $P_{\text{KS}} > 0.1$ in all cases. In what follows, we adopt this value as the boundary between distributions that are consistent and inconsistent with each other.

MGRSs (1)	[50, 1000] (2)	[50, 200] (3)	[200, 350] (4)	[350, 1000] (5)
1–2	0.70	0.60	0.60	0.97
1–3	0.44	0.26	0.99	0.69
1–4	0.82	0.98	0.46	0.31
2–3	0.36	0.11	0.58	0.81
2–4	0.73	0.65	0.42	0.26
3–4	0.40	0.48	0.62	0.36

constructed four independent MGRSs (all with $b_{\text{vel}} = 0.5$) from a set of independent simulation boxes (see Yang et al. 2004 for details). For each of these four MGRSs, we construct a group catalogue as described in Section 3.1 and compare the resulting $P(\mathcal{R})$ using the KS statistic. The results are listed in Table 2. In all cases, we find that $P_{\text{KS}} > 0.1$. In what follows, we therefore consider this probability value to indicate the boundary between two distributions being statistically equivalent or not.

The second issue that needs to be addressed concerns the number density of rich groups. As shown in YMBJ, our MGRSs overpredict the number of rich groups compared with the 2dFGRS. This implies that either the average mass-to-light ratio of groups and clusters is significantly higher than normally assumed, or that the power-spectrum normalization $\sigma_8 \simeq 0.75$ as opposed to 0.9 (cf. van den Bosch, Mo & Yang 2003; Yang et al. 2004; van den Bosch et al. 2005; YMBJ). As shown in Fig. 3, more massive groups are less severely affected by incompleteness and interlopers than less massive groups. Therefore, the \mathcal{R} distribution of all groups is sensitive to the relative fractions of groups of different masses (or different $\hat{\sigma}_8$). In order for the mismatch in group abundances between mock-up and data to not influence our results, we randomly remove massive groups from our MGRSs until we match the relative number of groups in the various $\hat{\sigma}_8$ bins.

In Fig. 4, we compare $P(< |\mathcal{R}|)$ obtained from our 2dFGRS group catalogue (grey dots), to those obtained from three MGRSs with different values of b_{vel} , as indicated. The upper-left panel plots the results using all groups in the full range of $\hat{\sigma}_8$ considered. The MGRS with $b_{\text{vel}} = 0.0$ (i.e. the one that fulfills the null hypothesis of the CGP) seems to be in reasonable agreement with the 2dFGRS data, although its $P(|\mathcal{R}|)$ appears to be somewhat narrower. Nevertheless, it clearly fits the data better than the MGRSs with $b_{\text{vel}} = 0.5$ or $b_{\text{vel}} = 1.0$. The other three panels of Fig. 4 plot $P(< |\mathcal{R}|)$ for three different bins of $\hat{\sigma}_8$, as indicated. Although the results are somewhat more noisy (because of the smaller number of groups involved), a comparison with the various MGRSs also suggests best-fitting values for b_{vel} in the range 0–0.5.

In order to make the comparison between 2dFGRS and MGRS more quantitative, Fig. 5 plots the logarithm of the KS probability, P_{KS} , that the $P(|\mathcal{R}|)$ of the 2dFGRS and the MGRS are drawn from the same distribution. Results are shown as function of b_{vel} and for the same four intervals of $\hat{\sigma}_8$ as in Fig. 4 (see also Table 1). When all groups with $50 \text{ km s}^{-1} \leq \hat{\sigma}_8 \leq 1000 \text{ km s}^{-1}$ are considered, the

data are consistent with the null hypothesis of the CGP. However, the data are equally consistent with a small but non-zero velocity bias, while models with $b_{\text{vel}} > 0.3$ are strongly ruled out by the data. When analysing the three $\hat{\sigma}_8$ subsamples, one notices a weak trend: more massive haloes seem to require a higher value of b_{vel} . If we adopt $P_{\text{KS}} > 0.1$ to indicate consistency with the data (see above), we obtain the following constraints on the velocity bias of the central galaxies: $b_{\text{vel}} \lesssim 0.4$ for $50 \text{ km s}^{-1} \leq \hat{\sigma}_8 \leq 200 \text{ km s}^{-1}$, $0.1 \lesssim b_{\text{vel}} \lesssim 0.5$ for $200 \text{ km s}^{-1} \leq \hat{\sigma}_8 \leq 350 \text{ km s}^{-1}$ and $0.2 \lesssim b_{\text{vel}} \lesssim 0.6$ for $350 \text{ km s}^{-1} \leq \hat{\sigma}_8 \leq 1000 \text{ km s}^{-1}$.

Finally, we address the issue of velocity bias for the *satellite* galaxies. When constructing our MGRSs, we have made the assumption that satellite galaxies have the same spatial and kinematic distribution as dark matter particles. These assumptions are, at least partially, supported by observations of cluster galaxies (e.g. Carlberg et al. 1997; van der Marel et al. 2000; Lin, Mohr & Stanford 2004). On the other hand, numerical simulations seem to suggest that satellite galaxies may have a small velocity bias: Using smoothed particle hydrodynamics (SPH) simulations, Berlind et al. (2003) found that satellites in intermediate-mass haloes are slightly colder than the dark matter in their host haloes, corresponding to a velocity bias of $b_{\text{sat}} \equiv \sigma_{\text{sat}}/\sigma_{\text{DM}} \sim 0.9$. On the other hand, Faltenbacher et al. (2005), using gas dynamical simulations of clusters, found that galaxies are slightly hotter than the dark matter ($b_{\text{sat}} \sim 1.1$), in good agreement with results obtained for dark matter subhaloes *without* gas (see Diemand, Moore & Stadel 2004 and references therein). Because both \bar{v}_s and $\hat{\sigma}_s$ scale linearly with b_{sat} , but the velocity of the brightest halo galaxy does not, a velocity bias $b_{\text{sat}} \neq 1$ will have a small effect on the measure of \mathcal{R} . To test the magnitude of this effect, we constructed two MGRSs with $b_{\text{vel}} = 0.5$: one with $b_{\text{sat}} = 0.9$ and the other with $b_{\text{sat}} = 1.1$. The \mathcal{R} distributions of these MGRSs are shown in Fig. 6. As expected, a lower value of b_{sat} results in a somewhat broader distribution of \mathcal{R} values. To a good approximation, reducing (increasing) b_{sat} by 0.1 has a similar effect on $P(\mathcal{R})$ as an increase (decrease) of b_{vel} by a similar value. Using this simple rule of thumb, the reader can convert the constraints on b_{vel} obtained here, which correspond to $b_{\text{sat}} = 1.0$, to their own favorite value for the velocity bias of satellite galaxies.

7 DISCUSSION

Our finding that the brightest halo galaxies in massive groups and clusters have, on average, a non-zero specific kinetic energy has two possible interpretations. First of all, the central galaxy may not be at rest with respect to the *virialized* dark matter halo. This scenario, which we hereafter refer to as the Non-Relaxed Galaxy (NRG) scenario, is illustrated in the left-hand panel of Fig. 7. The right-hand panel depicts the second possible scenario: that of a Non-Relaxed Halo (NRH). In this case, the brightest halo galaxy is located at rest with respect to the *minimum* of the dark matter potential, but the dark matter mass distribution is not relaxed and reveals a clear $m = 1$ mode (i.e. the potential minimum does not coincide with the barycentre). In both scenarios, the brightest halo galaxy has a net velocity with respect to the coordinate frame defined by the mean motion of the satellites. Note that, although our MGRSs are based on the NRG scenario, to first order it also mimics the NRH scenario, so that our comparison between MGRS and 2dFGRS applies to both cases.

Of the two scenarios illustrated in Fig. 7, the most likely one is the NRH scenario. It seems to fit naturally within a hierarchical picture of structure formation, where haloes continue to grow in mass by accretion and merging. It is also in accord with our finding

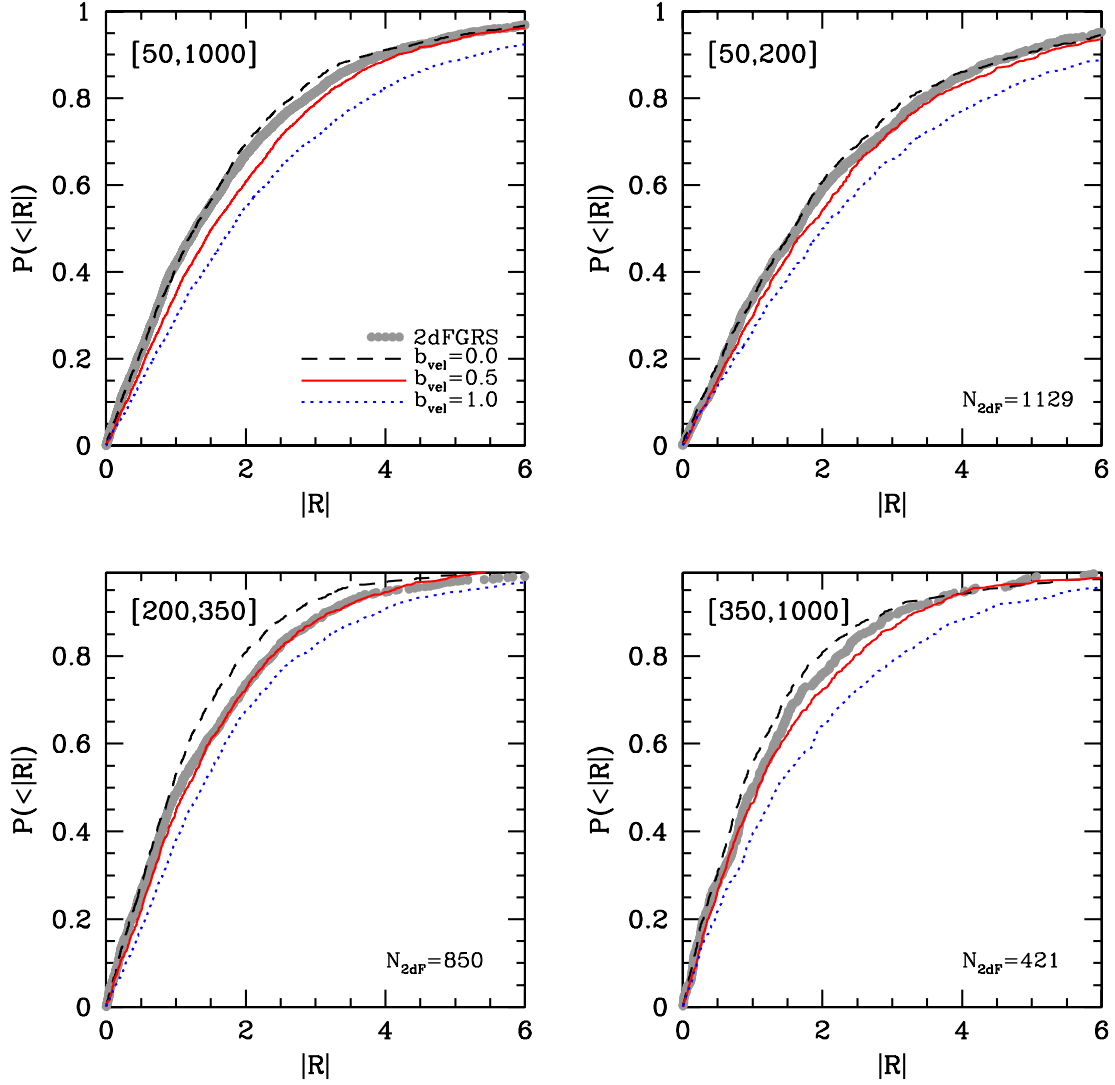


Figure 4. The cumulative distributions of $|R|$ obtained from the groups in the 2dFGRS (grey dots), compared with those obtained from three of our MGRSs, which only differ in their value of b_{vel} , as indicated in the upper-left panel. Results are shown for four intervals in $\hat{\sigma}_s$, indicated in square brackets in each panel. The number of 2dF groups in each of the three subsamples is indicated.

that b_{vel} is larger in more massive haloes, which form later and are thus expected to be less relaxed. Using a combination of numerical simulations and semi-analytical models of galaxy formation in a Λ CDM cosmology, Diaferio et al. (1999) found that the central galaxy has an average velocity with respect to the halo barycentre of $\sim 80 \text{ km s}^{-1}$. With a median halo mass of $\sim 10^{13} h^{-1} M_{\odot}$, this corresponds to $\langle |v_{\text{cen}}| \rangle = 0.23 \sqrt{\langle |v_{\text{DM}}|^2 \rangle}$.³ A similar result was obtained by Yoshikawa, Jing & Börner (2003) who, using an SPH simulation of galaxy formation in a Λ CDM universe, found that $\langle |v_{\text{cen}}| \rangle = 0.27 \sqrt{\langle |v_{\text{DM}}|^2 \rangle}$. Finally, Berlind et al. (2003), using similar SPH simulations, found that $\langle |v_{\text{cen}}| \rangle \simeq 0.2 \langle |v_{\text{DM}}| \rangle$. If we make the simplifying assumption that

$$b_{\text{vel}} = \frac{\sqrt{\langle |v_{\text{cen}}|^2 \rangle}}{\sqrt{\langle |v_{\text{DM}}|^2 \rangle}} \simeq \frac{\langle |v_{\text{cen}}| \rangle}{\sqrt{\langle |v_{\text{DM}}|^2 \rangle}} \simeq \frac{\langle |v_{\text{cen}}| \rangle}{\langle |v_{\text{DM}}| \rangle} \quad (19)$$

³ Here, we have used equation (8) and the assumption of isotropy to relate halo mass to the three-dimensional velocity dispersion of the dark matter particles.

(where the approximations are typically accurate to a few per cent), these results imply values for b_{vel} in the range 0.2 to 0.30. A comparison with Fig. 5 shows that these results are in good overall agreement with the constraints on b_{vel} obtained from our 2dFGRS group catalogue.

Unfortunately, Yoshikawa et al. (2003) did not investigate whether the most massive halo galaxy has a net velocity with respect to the most bound halo particle, while in Diaferio et al. (1999) the central galaxy is associated with the most bound halo particles by construction. Therefore, we can not use either of these results to discriminate between the NRG and NRH scenarios. Berlind et al. (2003), however, showed that, in their SPH simulation, central galaxies are closer to the most bound dark matter particle of their halo (mean separation of ~ 2 per cent of the virial radius) than to the centre of mass of their halo (mean separation of ~ 10 per cent of the virial radius). This clearly points towards the NRH scenario. We thus conclude that the simulation results are in good overall agreement with the results obtained here, and that they seem to support the NRH scenario as the most likely cause for the offset between the velocity of the central galaxy and that of the satellites.

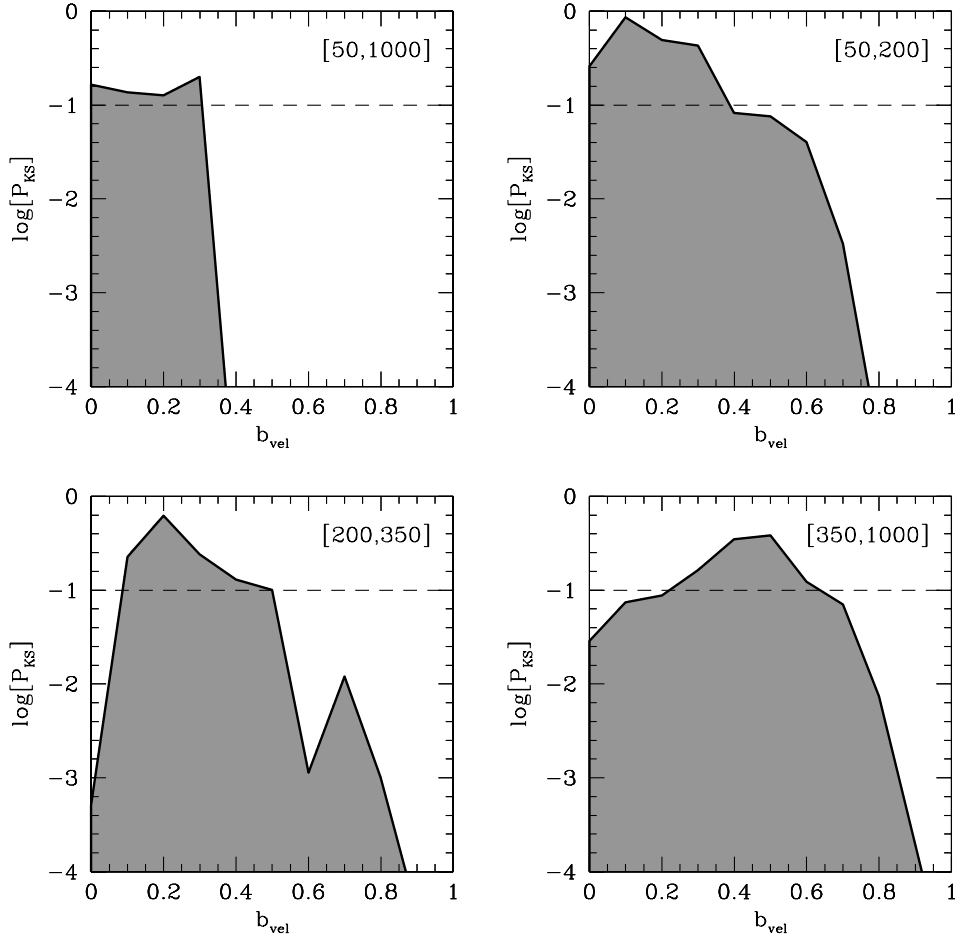


Figure 5. The KS probability that the \mathcal{R} distribution obtained from the 2dFGRS groups is consistent with that obtained from our MGRSs, as function of b_{vel} . Results are shown for four σ_s intervals, indicated in square brackets in each panel. The horizontal, dashed line in each panel indicates $P_{\text{KS}} = 0.1$: based on estimates of the scatter due to cosmic variance, we consider two distributions to be statistically equivalent when $P_{\text{KS}} > 0.1$. Whereas low-mass groups are consistent with $b_{\text{vel}} = 0$, the brightest galaxies in more massive groups are inconsistent with the null hypothesis of the CGP.

The NRG scenario appears unlikely at first sight, as dynamical friction against the highly concentrated dark matter halo probably quickly damps any oscillatory motion. On the other hand, if dark matter haloes are cored, rather than cusped, the oscillations may persist for a much longer time (cf. Bontekoe 1988). This possibility is interesting in the light of various independent claims for cored dark matter haloes, based on rotation curves of dwarf and low surface brightness galaxies (e.g. Flores & Primack 1994; Moore 1994; de Blok et al. 2001; Borriello & Salucci 2001; de Blok & Bosma 2002, but see also van den Bosch et al. 1999; van den Bosch & Swaters 2001; Dutton et al. 2005), on the observed pattern speeds of barred galaxies (Debattista & Sellwood 1998, 2000), and on the longevity of the lopsidedness of disc galaxies (Levine & Sparke 1998). A more in-depth study of the damping rate of these kinds of oscillations in dark matter haloes, both cusped and cored, could shed more light on these issues.

Independent of which of the aforementioned scenarios is responsible for the non-zero velocity bias of the brightest halo galaxies, it has important implications for various areas in astrophysics. First of all, it has an important impact on the use of satellite kinematics to infer halo masses. Because the number of detectable satellites in individual systems is generally small, one typically stacks the data

on many host–satellite pairs to obtain *statistical* estimates of halo masses (Erickson, Gottesman & Hunter 1987; Zaritsky et al. 1993; Zaritsky & White 1994; Zaritsky et al. 1997; McKay et al. 2002; Brainerd & Specian 2003; Prada et al. 2003; van den Bosch et al. 2004). The halo mass is typically derived from the dispersion, σ_{cs} , of the distribution of the velocity difference between host and satellite galaxies. This derivation rests on the (standard) assumptions that the host galaxies (i.e. the brightest halo galaxies) are at rest with respect to the centre of a *relaxed* dark matter halo. If the satellite galaxies have the same kinematics as dark matter particles, then $\sigma_{\text{cs}} = \sigma_{\text{dm}} \propto M^{1/3}$. However, in the case of the NRH scenario, one simply can not use (satellite) kinematics to infer reliable halo masses, as the crucial assumption of a virialized system is not correct. In the case of the NRG scenario, on the other hand, the system is relaxed but, because of the non-zero velocity bias, we have that $\sigma_{\text{cs}} = \sqrt{1 + b_{\text{vel}}} \sigma_{\text{dm}}$. If one were not to correct for b_{vel} , the inferred halo mass would be overestimated by a factor $(1 + b_{\text{vel}})^{3/2}$ (corresponding to ~ 1.5 in the case of $b_{\text{vel}} = 0.3$).

The results presented here also have potentially important implications for (strong) gravitational lensing. In both the NRG and NRH scenarios, one expects a strong, ‘external’ shear due to the dark matter halo, which should leave signatures in the image

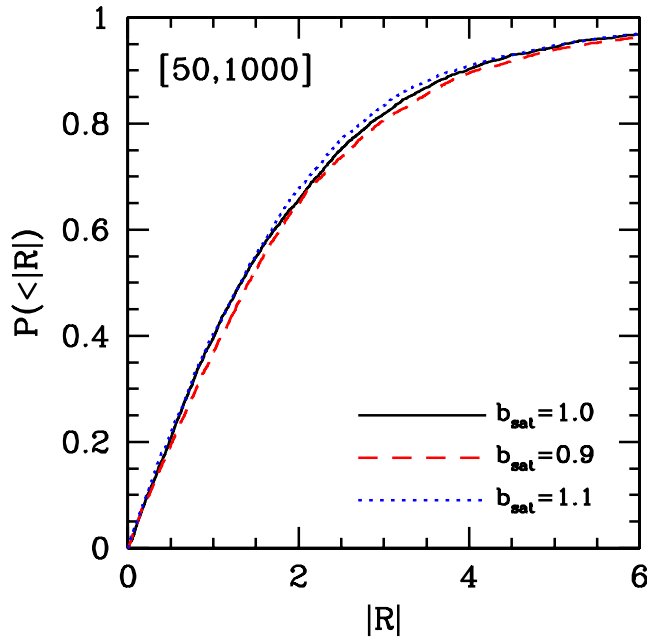


Figure 6. The cumulative distributions of $|\mathcal{R}|$ for three MGRSs with $b_{\text{vel}} = 0.5$, but with different values for the velocity bias b_{sat} of the satellite galaxies as indicated. Note that $b_{\text{sat}} = 1.0$ corresponds to no velocity bias (i.e. satellite galaxies have the same velocity distribution as dark matter particles). Overall the impact of the velocity bias of the satellite galaxies on $P(\mathcal{R})$ is small (see text for detailed discussion).

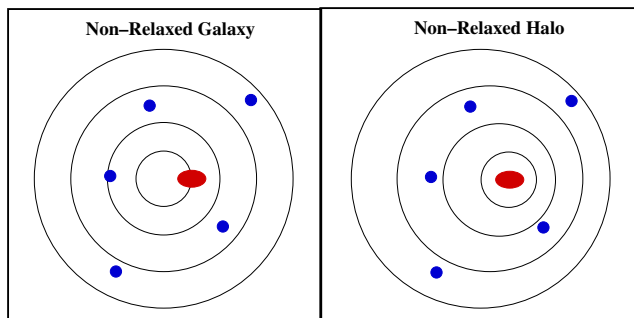


Figure 7. An illustration of the two different configurations that are both consistent with our inferred offset between the brightest halo galaxy and the satellite galaxies. Contours depict equipotentials of the dark matter haloes, while filled ellipses and circles represent the brightest halo galaxies and satellite galaxies, respectively. In the Non-Relaxed Galaxy (NRG) scenario, the brightest halo galaxy oscillates in a fully relaxed halo. In the Non-Relaxed Halo (NRH) scenario, on the other hand, the central galaxy coincides with the minimum of the halo potential, but the centres of different equipotential surfaces are offset from each other. See text for a detailed discussion.

configurations and time delays of the lens. In fact, this ‘external’ shear may already have been detected. As shown in Keeton, Kochanek & Seljak (1997), fitting four-image lenses almost always requires an independent external shear that is not aligned with the light of the lens. Although this may reflect a misalignment between the luminous galaxy and dark matter halo, in agreement with the results presented here, there are alternative sources of external shear (nearby galaxies, large-scale structure along the line-of-sight) that may leave a similar signal in the lens configuration. A more thorough, systematic study of multiple-lensed systems is therefore re-

quired to put constraints on the spatial bias of the brightest halo galaxies. In fact, strong gravitational lensing is probably the only method that can be used to detect an offset between halo and galaxy in *individual* systems, and to discriminate between the NRH and NRG scenarios.

A non-zero $\langle r_{\text{cen}} \rangle$ also impacts on the internal structure and dynamics of central galaxies. As the galaxy oscillates in the dark matter halo (NRG scenario), or the halo relaxes around the central galaxy (NRH scenario), it is constantly subjected to tidal forces that may trigger bar instabilities in otherwise stable discs, may cause excessive heating of the disc and may create lopsidedness (Levine & Sparke 1998; Noordermeer, Sparke & Levine 2001). Detailed studies have revealed lopsidedness (either in the kinematics or the photometry) in about half of all disc galaxies studied (e.g. Richter & Sancisi 1994; Zaritsky & Rix 1997; Haynes et al. 1998; Matthews, van Driel & Gallagher 1998; Rudnick & Rix 1998; Swaters 1999). In fact, as shown by Bissantz, Englmaier & Gerhard (2003), the morphology and kinematics of gas in the inner few kpc of the Milky Way (in particular the 3 kpc arm) may indicate the presence of a similar $m = 1$ asymmetry in our own galaxy (cf. Fux 1999). The high frequency of lopsided and barred disc galaxies therefore seems to be in support of a non-zero $\langle r_{\text{cen}} \rangle$, whether it reflects a non-relaxed halo or a non-relaxed galaxy. Taking our results at face value, it is clear that any study of disc stability that ignores these strong distortions and time variability of the potential may be missing an essential ingredient.

Finally, as mentioned in the introduction, a non-zero b_{vel} also plays a role in halo occupation models. Using the method described in detail in Yang et al. (2004), we computed the projected two-point correlation function and pairwise peculiar velocity dispersions of MGRSs $M_{0.0}$, $M_{0.5}$ and $M_{1.0}$. The differences are found to be extremely small, well below the errors due to cosmic variance. Therefore, for all practical purposes, it suffices to model the phase-space parameters of galaxies in dark matter haloes with $b_{\text{vel}} = 0$ (as is generally done), when computing galaxy–galaxy correlation functions based on halo occupation distributions. Furthermore, as shown in Yang et al. (2005d), the effect of a non-zero b_{vel} is even small in the galaxy–dark matter cross-correlation function.

8 CONCLUSIONS

According to the standard paradigm of structure formation, the brightest galaxy in a dark matter halo should reside at rest at the centre of the potential well. In order to test this CGP, we used the halo-based galaxy group finder YMBJ to construct group catalogues from the 2dFGRS and SDSS. For each group, we compute the statistic \mathcal{R} , defined as the difference between the velocity of the brightest group galaxy and the average velocity of the other group members (satellites), normalized by the unbiased estimator of the velocity dispersion of the satellite galaxies. If the null hypothesis of the CGP is correct, \mathcal{R} should follow a Student t distribution. If, on the other hand, the brightest halo galaxies have a non-zero velocity bias with respect to the satellite galaxies, the \mathcal{R} distribution should be significantly broader. The applicability of this ‘ \mathcal{R} test’ depends critically on how well one can group those galaxies that belong to the same dark matter halo. Although our group finder is well tested and calibrated, it is not perfect and unavoidably selects interloper galaxies as group members. In addition, redshift surveys suffer from various incompleteness effects. We have shown that these effects result in a broadening of the \mathcal{R} distribution, which, when not accounted for, may give the false impression that the CGP is ruled out.

In order to take interlopers and incompleteness effects into account, and thus allow for a fair comparison with the data, we construct detailed MGRSs that can be compared with the 2dFGRS on a one-to-one basis. We apply our \mathcal{R} statistic to the galaxy groups selected from these MGRS, which we compare to those obtained from the 2dFGRS using the KS test. This shows that the CGP is only consistent with the data for haloes with $\hat{\sigma}_s \lesssim 200 \text{ km s}^{-1}$ (corresponding to haloes with $M \lesssim 10^{13} h^{-1} M_\odot$). In more massive haloes, however, we find a clear indication that the brightest halo galaxies have a non-zero specific kinetic energy. In massive clusters, this specific kinetic energy may be as large as ~ 25 per cent of that of the satellite galaxies (corresponding to $b_{\text{vel}} = 0.5$). Although this may seem large, for a typical, relaxed, CDM halo it only corresponds to an expectation value for the offset between the galaxy and halo of ~ 3 per cent of the virial radius, comparable to the characteristic radius of the galaxy.

We have focussed mainly on the \mathcal{R} distributions obtained from the 2dFGRS, simply because we have accurate MGRSs available for this data set. However, we have shown that the \mathcal{R} distributions obtained from groups in the SDSS are in excellent agreement with those obtained from the 2dFGRS, suggesting that the SDSS is also inconsistent with the CGP.

What is the origin of this offset between the central galaxy and the rest-frame defined by the satellites? Probably the most likely explanation is that the majority of dark matter haloes are not yet fully relaxed. In this case, the brightest halo galaxy may still coincide with the *minimum* of the potential well, but that minimum does not coincide with the centre of mass measured over the entire halo. This picture is consistent with our finding that the specific kinetic energy of the brightest halo galaxies is larger in more massive haloes, which form later, and also has support from numerical simulations, which reveal a velocity bias of the brightest halo galaxies that is very similar to that found here (Diaferio et al. 1999; Yoshikawa et al. 2003). In particular, the simulations of Berlind et al. (2003) have shown that the brightest halo galaxy is, on average, much closer to the most bound particle than it is to the halo centre of mass.

An alternative explanation for the non-zero velocity of the brightest halo galaxies with respect to the satellite galaxies may be that the halo is relaxed, but that the brightest halo galaxy oscillates in the central potential well. If the dark matter halo is strongly concentrated, as expected for typical CDM haloes, one would naively expect that any such oscillation is quickly damped by dynamical friction. However, this damping time-scale may be significantly longer if there is less dark matter in the centre of the halo than anticipated; i.e. if the density distribution is cored rather than cusped. This possibility is interesting in light of recent claims for cored haloes based on the observed rotation curves and bar pattern speeds of disc galaxies.

In either case, the brightest halo galaxy is expected to experience a time-varying tidal field. This strongly questions the applicability of (numerical) studies of galaxy dynamics and in particular of stability analyses that make the assumption that the galaxy is at rest at the centre of a relaxed dark matter halo. In particular, it may explain the high frequency and longevity of bars and lopsidedness in disc galaxies. The fact that we find evidence for a non-zero velocity of the brightest halo galaxy with respect to the satellite galaxies also has important implications for the determination of halo masses based on the kinematics of host–satellite systems and for the modelling of strong gravitational lenses. For the purpose of computing galaxy–galaxy correlation functions based on halo occupation models, however, one can safely ignore the fact that the CGP does not

hold and make the simple Ansatz that the brightest halo galaxy resides at rest at the halo centre.

ACKNOWLEDGMENTS

We are grateful to Michael Blanton for his help with the NYU-VAGC, to the anonymous referee for a detailed report that greatly helped to improve the paper, and to Victor Debattista, Savvas Koushiappas, George Lake, Shude Mao, Ben Moore, Peter Schneider, Joachim Stadel and Simon White for useful discussions.

REFERENCES

- Abazajian K. et al., 2004, *AJ*, 128, 502
- Berlind A. A., Weinberg D. H., 2002, *ApJ*, 575, 587
- Berlind A. A. et al., 2003, *ApJ*, 593, 1
- Binney J. J., Tremaine S. D., 1987, *Galactic Dynamics*. Princeton Univ. Press, Princeton
- Bissantz N., Englmaier P., Gerhard O. E., 2003, *MNRAS*, 340, 949
- Blanton M. R. et al., 2005, *AJ*, 129, 2562
- de Blok W. J. G., Bosma A., 2002, *A&A*, 385, 816
- de Blok W. J. G., McGaugh S. S., Bosma A., Rubin V. C., 2001, *ApJ*, 552, L23
- Bontekoe T. R., 1988, PhD thesis, Univ. Groningen
- Borriello A., Salucci P., 2001, *MNRAS*, 323, 285
- Brainerd T. G., Specian M. A., 2003, *ApJ*, 593, L7
- Carlberg R. G. et al., 1997, *ApJ*, 485, L13
- Cole S., Lacey C. G., Baugh C. M., Frenk C. S., 2000, *MNRAS*, 319, 168
- Cole S. et al., 2001, *MNRAS*, 326, 255
- Colless M., et al., 2001, *MNRAS*, 328, 1039
- Debattista V. P., Sellwood J. A., 1998, *ApJ*, 493, L5
- Debattista V. P., Sellwood J. A., 2000, *ApJ*, 543, 704
- Diaferio A., Kauffmann G., Colberg J. M., White S. D. M., 1999, *MNRAS*, 307, 537
- Diemand J., Moore B., Stadel J., 2004, *MNRAS*, 352, 535
- Dutton A. A., Courteau S., de Jong R. S., Carignan C., 2005, *ApJ*, 619, 218
- Eke V. R., Navarro J. F., Steinmetz M., 2001, *ApJ*, 554, 114
- Erickson L. K., Gottesman S. T., Hunter J. H., 1987, *Nat*, 325, 779
- Faltenbacher A., Kravtsov A. V., Nagai D., Gottlöber S., 2005, *MNRAS*, 358, 139
- Flores R. A., Primack J. R., 1994, *ApJ*, 427, L1
- Fux R., 1999, *A&A*, 345, 787
- Hawkins E. et al., 2003, *MNRAS*, 346, 78
- Haynes M. P., Hogg D. E., Maddalena R. J., Roberts M. S., van Zee L., 1998, *AJ*, 115, 62
- Hernquist L., 1990, *ApJ*, 356, 359
- Hill J. M., Hintzen P., Oegerle W. R., Romanishin W., Lesser M. P., Eisenhamer J. D., Batuski D. J., 1988, *ApJ*, 332, L23
- Jing Y. P., Suto Y., 2002, *ApJ*, 574, 538
- Jones C., Mandel E., Schwarz J., Forman W., Murray S. S., Harnden F. R., 1979, *ApJ*, 234, L21
- Kauffmann G., White S. D. M., Guiderdoni B., 1993, *MNRAS*, 264, 201
- Keeton C. R., Kochanek C. S., Seljak U., 1997, *ApJ*, 482, 604
- Levine S. E., Sparke L. S., 1998, *ApJ*, 496, L13
- Lin Y.-T., Mohr J. J., Stanford S. A., 2004, *ApJ*, 610, 745
- McKay T. A. et al., 2002, *ApJ*, 571, L85
- Madgwick D. S. et al., 2002, *MNRAS*, 333, 133
- Magliocchetti M., Porciani C., 2003, *MNRAS*, 346, 186
- van der Marel R. P., Magorrian J., Carlberg R. G., Yee H. K. C., Ellingson E., 2000, *AJ*, 119, 2038
- Matthews L. D., van Driel W., Gallagher J. S., 1998, *AJ*, 116, 1169
- Moore B., 1994, *Nat*, 370, 629
- Mulchaey J. S., Zabludoff A. I., 1998, *ApJ*, 496, 73
- Navarro J. F., Frenk C. S., White S. D. M., 1997, *ApJ*, 490, 493
- Noordermeer E., Sparke L. S., Levine S. E., 2001, *MNRAS*, 328, 1064
- Norberg P. et al., 2002, *MNRAS*, 336, 907

- Oegerle W. R., Hill J. M., 1994, *AJ*, 107, 857
- Oegerle W. R., Hill J. M., 2001, *AJ*, 122, 2858
- Postman M., Lubin L. M., Gunn J. E., Oke J. B., Hoessel J. G., Schneider D. P., Christensen J. A., 1996, *AJ*, 111, 615
- Prada F. et al., 2003, *ApJ*, 598, 260
- Quintana H., Lawrie D. G., 1982, *AJ*, 87, 1
- Richter O.-G., Sancisi R., 1994, *A&A*, 290, L9
- Rudnick G., Rix H.-W., 1998, *AJ*, 116, 1163
- Scoccimarro R., Sheth R. K., Hui L., Jain B., 2001, *ApJ*, 546, 20
- Sharples R. M., Ellis R. S., Gray P. M., 1988, *MNRAS*, 231, 479
- Sheth R. K., Mo H. J., Tormen G., 2001, *MNRAS*, 323, 1
- Somerville R. S., Primack J. R., 1999, *MNRAS*, 310, 1087
- Swaters R. A., 1999, PhD thesis, Univ. Groningen
- Tinker J., Weinberg D. H., Zheng Z., Zehavi I., 2004, *ApJ*, submitted, (astro-ph/0411777)
- van den Bosch F. C., Swaters R. A., 2001, *MNRAS*, 325, 1017
- van den Bosch F. C., Robertson B. E., Dalcanton J. J., de Blok W. J. G., 1999, *AJ*, 119, 1579
- van den Bosch F. C., Yang X., Mo H. J., 2003, *MNRAS*, 340, 771
- van den Bosch F. C., Mo H. J., Yang X., 2003, *MNRAS*, 345, 923
- van den Bosch F. C., Norberg P., Mo H. J., Yang X., 2004, *MNRAS*, 352, 1302
- van den Bosch F. C., Mo H. J., Yang X., Norberg P., 2005, *MNRAS*, 356, 1233
- White S. D. M., Rees M. J., 1978, *MNRAS*, 183, 341
- Yang X., Mo H. J., van den Bosch F. C., 2003, *MNRAS*, 339, 1057
- Yang X., Mo H. J., Jing Y. P., van den Bosch F. C., Chu Y., 2004, *MNRAS*, 350, 1153
- Yang X., Mo H. J., van den Bosch F. C., Jing Y. P., 2005a, *MNRAS*, 356, 1293 (YMBJ)
- Yang X., Mo H. J., van den Bosch F. C., Jing Y. P., 2005b, *MNRAS*, 357, 608
- Yang X., Mo H. J., Jing Y. P., van den Bosch F. C., 2005c, *MNRAS*, 358, 217
- Yang X., Mo H. J., van den Bosch F. C., Weinmann S. M., Li C., Jing Y. P., 2005d, *MNRAS*, in press (doi:10.1111/j.1365-2966.2005.09351.x) (astro-ph/0504477)
- York D. et al., 2000, *AJ*, 120, 1579
- Yoshikawa K., Jing Y. P., Börner G., 2003, *ApJ*, 590, 654
- Zabludoff A. I., Mulchaey J. S., 1998, *ApJ*, 496, 39
- Zabludoff A. I., Huchra J. P., Geller M. J., 1990, *ApJS*, 74, 1
- Zaritsky D., Rix H. W., 1997, *ApJ*, 477, 118
- Zaritsky D., White S. D. M., 1994, *ApJ*, 435, 599
- Zaritsky D., Smith R., Frenk C. S., White S. D. M., 1993, *ApJ*, 405, 464
- Zaritsky D., Smith R., Frenk C. S., White S. D. M., 1997, *ApJ*, 478, 39
- Zehavi I. et al., 2004, preprint (astro-ph/0408569)
- Zheng Z. et al., 2004, preprint (astro-ph/0408564)

This paper has been typeset from a \LaTeX file prepared by the author.